

Summary

- **Box shaped** confidence regions for θ based on Bonferronized z-tests or -t-tests based on estimates $\hat{\theta}_j$

$$R(\mathbf{X}) = \{\theta \mid \hat{\theta}_j - \tilde{K}_\alpha \times \hat{\sigma}_{\hat{\theta}_j} \leq \theta_j \leq \hat{\theta}_j + \tilde{K}_\alpha \times \hat{\sigma}_{\hat{\theta}_j}, j=1, \dots, q\},$$

with

$$\tilde{K}_\alpha = t_{f_e}(\alpha'/2) \text{ or } \tilde{K}_\alpha = z(\alpha'/2), \alpha' = \alpha/q$$

$\hat{\sigma}_{\hat{\theta}_j}$ is the estimated standard error of $\hat{\theta}_j$.

The shape is determined by the values of $\hat{\sigma}_{\hat{\theta}_j}^2$, the *diagonal* elements of $\hat{V}[\hat{\theta}]$.

- **Ellipsoidal** confidence regions based on Hotelling's T^2 test:

$$R(\mathbf{X}) = \{\theta \mid (\theta - \hat{\theta})' \{\hat{V}[\hat{\theta}]\}^{-1} (\theta - \hat{\theta}) \leq K_\alpha^2\},$$

$$K_\alpha^2 = \chi_q^2(\alpha) \text{ or } \{(f_e q)/(f_e - q + 1)\} F_{q, f_e - q + 1}(\alpha).$$

The *shape* is determined by *eigenvalues* of $\hat{V}[\hat{\theta}]$. The *orientation* is determined by the *eigenvectors* of $\hat{V}[\hat{\theta}]$.

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Displays for Statistics 5401/8401

Lecture 13

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Class Web Page

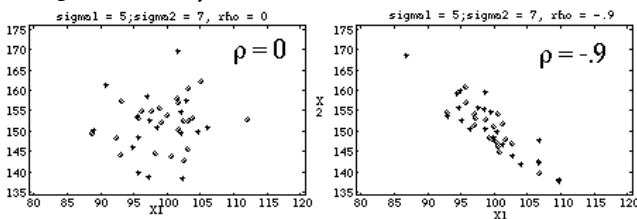
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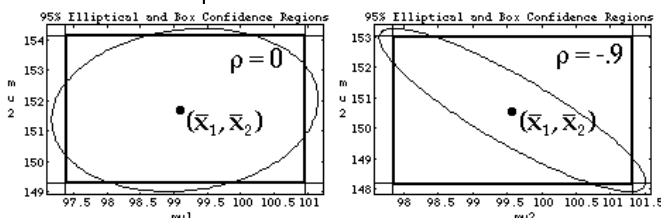
How far from "advertised" is a "Bonferoni box?" It depends on how correlated the data are. Here are two samples of 40, both from normal populations with

$$\sqrt{\sigma_{11}} = 5 \text{ and } \sqrt{\sigma_{22}} = 7.$$

The one on the left has $\rho = 0$; the one on the right has $\rho = -.9$



Here are 95% and 99% rectangular and elliptical confidence regions for μ based on these samples:



- The elliptical confidence regions have the same orientation as the "cloud of points" but are much smaller (compare the axis scales).
- For small ρ , there isn't a lot of difference between the elliptical and box regions: the box corners are slightly outside the ellipse; the ellipse ends and sides are slightly outside the box.
- When ρ is high, two corners of the box are distant from the ellipse and there is a lot of area outside the ellipse and in the box, and relatively little area outside the box and in the ellipse.
- As $\rho \rightarrow \pm 1$, actual confidence of box approaches $1 - \alpha/2$ instead of $1 - \alpha$ (97.5% instead of 95%, for example).

Simulation with $M = 10,000, n = 50$:

ρ	0	.9	.99
$1 - \hat{\alpha}$	0.9456	0.9614	0.9698

$1 - \hat{\alpha}$ estimates actual confidence level

- Both regions (box and ellipsoid) are centered at $\hat{\theta}$.
- The volume of a box shaped region is $\hat{\sigma}_{\hat{\theta}_1} \hat{\sigma}_{\hat{\theta}_2} \dots \hat{\sigma}_{\hat{\theta}_q} (2\tilde{K}_\alpha)^q$, $\tilde{K}_\alpha = z(\alpha'/2)$ or $t_{f_e}(\alpha'/2)$

For $q = 2$, area = $\hat{\sigma}_{\hat{\theta}_1} \hat{\sigma}_{\hat{\theta}_2} (2\tilde{K}_\alpha)^2$

- The volume of an ellipsoidal region is
$$\frac{\sqrt{\det(\hat{V}[\hat{\theta}])} \times \pi^{q/2} K_\alpha^q}{\Gamma((q+2)/2)}$$

When $q = 2$,

Area = $\sqrt{\{\det(\hat{V}[\hat{\theta}])\}} \pi K_\alpha^2$,
 $= \hat{\sigma}_{\hat{\theta}_1} \hat{\sigma}_{\hat{\theta}_2} \sqrt{\{1 - \hat{\rho}_{\hat{\theta}_1, \hat{\theta}_2}^2\}} K_\alpha^2 \pi$

$\hat{\rho}_{\hat{\theta}_1, \hat{\theta}_2}$ = estimated correlation between $\hat{\theta}_1$ and $\hat{\theta}_2$.

Note: For even $q = 2m$, $\Gamma((q+2)/2) = m!$

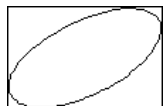
For odd $q = 2m-1$,

$\Gamma((q+2)/2) = 1 \times 3 \times \dots \times (2m-1) \sqrt{\pi/2^m}$

Bounding boxes for ellipsoids

Every ellipsoid has a rectangular *bounding box*.

- Each "face" (edge or wall) is perpendicular to a coordinate axis.
- Each face of the box is tangent to (touches at one point) the ellipsoid.



Bounding box when $p = 2$.

What is the size and shape of an ellipse's bounding box?

Note:

If Σ is a variance matrix, $\det(\Sigma)$ is the **generalized variance**, a single number which is sometimes used as a summary of how spread out a multivariate population is.

The volume of an ellipsoidal region is proportional to the square root of a generalized variance.

For fixed σ_{jj} , larger correlations result in smaller generalized variance.

Also for fixed $\text{trace}(\Sigma) = \sum_j \sigma_{jj} = \sum_j \lambda_j$, the more different are eigenvalues $\{\lambda_j\}$ of $\hat{V}[\hat{\theta}]$, the smaller is the generalized variance.

For instance, when $\lambda_1 = .55$ and $\lambda_2 = .45$, $\sqrt{\det(\Sigma)} = \sqrt{(.55 \times .45)} = 0.497$, while when $\lambda_1 = .9$ and $\lambda_2 = .1$, $\sqrt{\det(\Sigma)} = \sqrt{(.9 \times .1)} = 0.3 < 0.497$. In both cases $\lambda_1 + \lambda_2 = 1$.

Define E to be the inside and boundary of an ellipsoid with center at \mathbf{x}_0 . That is

$$E \equiv \{\mathbf{x} \mid (\mathbf{x} - \mathbf{x}_0)' \mathbf{Q}^{-1} (\mathbf{x} - \mathbf{x}_0) \leq K^2\}$$

where \mathbf{Q} is $q \times q$ symmetric positive definite ($\mathbf{Q} = \hat{V}(\hat{\theta})$ for a confidence ellipse).

Fact

The *bounding box* for E is the set

$$\{\mathbf{x} \mid x_{oj} - K\sqrt{q_{jj}} \leq x_j \leq x_{oj} + K\sqrt{q_{jj}}, j = 1, \dots, q\}$$

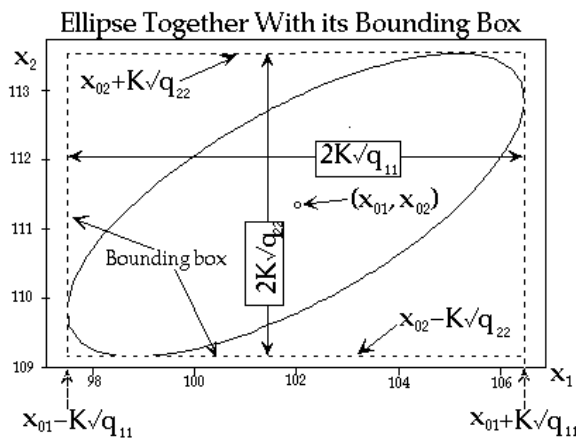
The bounding faces or planes come in parallel pairs, each pair perpendicular to a coordinate axis:

$$\{\mathbf{x} \mid x_j = x_{oj} - K\sqrt{q_{jj}}\} \text{ and } \{\mathbf{x} \mid x_j = x_{oj} + K\sqrt{q_{jj}}\}$$

These are perpendicular to the coordinate axis defined by

$$\mathbf{e}_j = [0 \ 0 \ \dots \ 0 \ 1 \ 0 \ \dots \ 0]'$$

and parallel the plane defined by containing the remaining axes $\{\mathbf{e}_i\}_{i \neq j}$.



When $p = 2$, the left and right tangent lines are the vertical lines defined by

$$x_1 = x_{01} - K\sqrt{q_{11}} \text{ and } x_1 = x_{01} + K\sqrt{q_{11}}$$

They are perpendicular to the x_1 axis.

The bottom and top tangents line are the horizontal lines defined by

$$x_2 = x_{02} - K\sqrt{q_{22}} \text{ and } x_2 = x_{02} + K\sqrt{q_{22}}$$

They are perpendicular to the x_2 axis.

This formulas of the bounding box are consequence of the following "fact":

- If \mathbf{x} is in E , every linear combination $\mathbf{l}'\mathbf{x} = \sum_i l_i x_i$ satisfies $\mathbf{l}'\mathbf{x}_0 - K\sqrt{(\mathbf{l}'\mathbf{Q}\mathbf{l})} \leq \mathbf{l}'\mathbf{x} \leq \mathbf{l}'\mathbf{x}_0 + K\sqrt{(\mathbf{l}'\mathbf{Q}\mathbf{l})}$
- Conversely, if these inequalities are true for every \mathbf{l} , then \mathbf{x} is in E

That is,

- When \mathbf{x} is *in* E , for every \mathbf{l} , the linear combination $\mathbf{l}'\mathbf{x}$ is *inside* the interval

$$\mathbf{l}'\mathbf{x}_0 \pm K\sqrt{(\mathbf{l}'\mathbf{Q}\mathbf{l})}$$

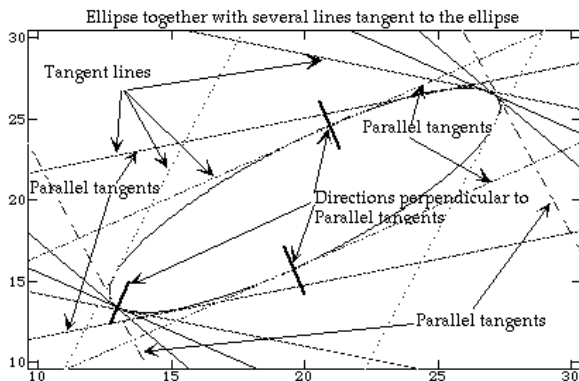
- When \mathbf{x} is *not* in E , there is some linear combination $\mathbf{l}'\mathbf{x}$ such that $\mathbf{l}'\mathbf{x} < \mathbf{l}'\mathbf{x}_0 - K\sqrt{(\mathbf{l}'\mathbf{Q}\mathbf{l})}$ (*outside* to left)

or

$$\mathbf{l}'\mathbf{x} > \mathbf{l}'\mathbf{x}_0 + K\sqrt{(\mathbf{l}'\mathbf{Q}\mathbf{l})} \text{ (*outside* to right)}$$

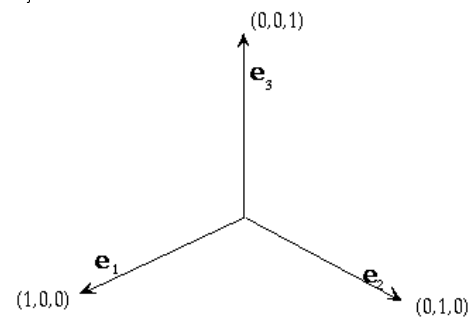
Note the use of \mathbf{Q} instead of \mathbf{Q}^{-1} here.

This description of an ellipsoid E corresponds to the obvious fact that the boundary and the interior of E consist exactly the points between *all* pairs of parallel tangent lines or planes.



The direction of each pair of lines is perpendicular to a vector \mathbf{l} (heavy lines in plot). Every vector \mathbf{l} determines two tangent lines (planes when $q > 2$).

A particular case is $\mathbf{l} = \mathbf{e}_j$, where, as before, \mathbf{e}_j is a "coordinate vector".



Then you have the particularly simple equations:

- $\mathbf{l}'\mathbf{x} = \sum_i l_i x_i = x_j$
- $\mathbf{l}'\mathbf{Q}\mathbf{l} = \sum_i \sum_j q_{ij} l_i l_j = q_{jj}$

When you apply the general result here you get the defining equations for the bounding box

$$\{\mathbf{x} \mid x_{0j} - K\sqrt{q_{jj}} \leq x_j \leq x_{0j} + K\sqrt{q_{jj}}, j = 1, \dots, q\}$$

Bounding boxes for ellipsoidal confidence regions

- A q-vector \mathbf{l} defines linear combinations $\mathbf{l}'\boldsymbol{\theta}$ and $\mathbf{l}'\hat{\boldsymbol{\theta}}$ of the elements of $\boldsymbol{\theta} = [\theta_1, \dots, \theta_q]'$ and $\hat{\boldsymbol{\theta}} = [\hat{\theta}_1, \dots, \hat{\theta}_q]'$.

- The estimated variance of $\mathbf{l}'\hat{\boldsymbol{\theta}}$ is

$$\hat{V}[\mathbf{l}'\hat{\boldsymbol{\theta}}] = \mathbf{l}'\hat{V}[\hat{\boldsymbol{\theta}}]\mathbf{l}$$

- The estimated *standard error* of $\mathbf{l}'\hat{\boldsymbol{\theta}}$ is

$$\hat{\sigma}_{\mathbf{l}'\hat{\boldsymbol{\theta}}} = \sqrt{\{\hat{V}[\mathbf{l}'\hat{\boldsymbol{\theta}}]\}} = \sqrt{\{\mathbf{l}'\hat{V}[\hat{\boldsymbol{\theta}}]\mathbf{l}\}}.$$

$\sqrt{\{\hat{V}[\mathbf{l}'\hat{\boldsymbol{\theta}}]\}}$ is $\sqrt{(\mathbf{l}'\mathbf{Q}\mathbf{l})}$ when $\mathbf{Q} = \hat{V}[\hat{\boldsymbol{\theta}}]$.

- The faces of the bounding box are at distances $K_\alpha \hat{\sigma}_{\hat{\theta}_j}$ from the center.

- The bounding box for the ellipsoid is **always bigger** than the Bonferroni box.

This means

$$P(R_{\text{Bounding box}}(\mathbf{X}) \text{ contains } \boldsymbol{\theta}) > P(R_{\text{Bonferroni box}}(\mathbf{X}) \text{ contains } \boldsymbol{\theta})$$

Since

$$P(R_{\text{Bonferroni box}}(\mathbf{X}) \text{ contains } \boldsymbol{\theta}) > 1 - \alpha,$$

the bounding box can be considered a $1 - \alpha$ confidence region for $\boldsymbol{\theta}$, but is *very conservative* with actual confidence level

$$P(R_{\text{Bounding box}}(\mathbf{X}) \text{ contains } \boldsymbol{\theta}) \gg 1 - \alpha$$

If

$$R(\mathbf{X}) = \{\boldsymbol{\theta} \mid (\boldsymbol{\theta} - \hat{\boldsymbol{\theta}})' \hat{V}[\hat{\boldsymbol{\theta}}]^{-1} (\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}) \leq K_\alpha^2\},$$

is an ellipsoidal confidence region then its bounding box is centered at $\hat{\boldsymbol{\theta}}$ with edges parallel with lengths $2K_\alpha \sqrt{\hat{\sigma}_{\hat{\theta}_j}}$.

As usual

$$K_\alpha = \chi_q(\alpha) = \sqrt{\{\chi_q^2(\alpha)\}} \text{ (large sample)}$$

or

$$K_\alpha = \sqrt{\{(qf_e/(f_e - q + 1))F_{q, f_e - q + 1}(\alpha)\}} \text{ (small)}$$

- This is *the same shape* -- but larger -- as the box-shaped confidence region obtained by Bonferronizing separate tests of each θ_j .

The lengths of the sides of the "Bonferroni" box are

$$2 \times z(\alpha'/2) \hat{\sigma}_{\hat{\theta}_j} \text{ or } 2 \times t_{f_e}(\alpha'/2) \hat{\sigma}_{\hat{\theta}_j}$$

$$\alpha' = \alpha/q.$$

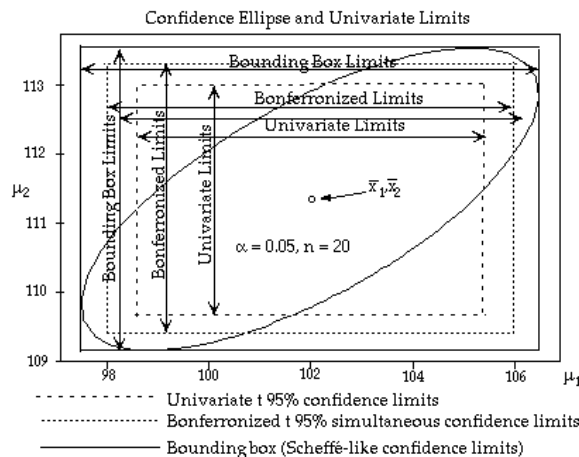
q = number of parameters in $\boldsymbol{\theta}$, and might not be p = number of variables.

The sides of the bounding box define simultaneous confidence limits for the parameters, since

$$P(\hat{\theta}_j - K \hat{\sigma}_{\hat{\theta}_j} \leq \theta_j \leq \hat{\theta}_j + K \hat{\sigma}_{\hat{\theta}_j}, j = 1, \dots, q) =$$

$$P(R_{\text{bounding box}}(\mathbf{X}) \text{ contains } \boldsymbol{\theta}) \gg 1 - \alpha$$

Generally, these are *very conservative* confidence limits (confidence $\gg 1 - \alpha$).



Here $q = 2$, $\boldsymbol{\theta} = \boldsymbol{\mu}$, $\hat{\boldsymbol{\theta}} = \bar{\mathbf{x}}$

I continued the simulation reported on before to estimate the actual confidence level of simultaneous confidence limits based on the bounding box.

Estimated Confidence Levels

ρ	0	.9	.99
$1 - \hat{\alpha}$.9702	0.9776	0.9848

These are unacceptably larger than the intended confidence $1 - \alpha = .95$.

Vocabulary

I refer to bounding box limits and their generalization to linear combinations of parameters as **ellipsoidal limits**.

With $\hat{\theta} = \bar{x}$, and $\hat{V}[\hat{\theta}] = (1/n)S$,

- $\mathbf{l}_{ij}'\hat{\theta} = \bar{x}_i - \bar{x}_j$
- $\hat{\sigma}_{\mathbf{l}_{ij}'\hat{\theta}} = \hat{\sigma}_{\bar{x}_i - \bar{x}_j} = \sqrt{\{\mathbf{l}_{ij}'\hat{V}[\hat{\theta}]\mathbf{l}_{ij}\}}$
 $= \sqrt{\{\hat{v}_{ii} - 2\hat{v}_{ij} + \hat{v}_{jj}\}} = \sqrt{\{(1/n)(s_{ii} - 2s_{ij} + s_{jj})\}}$

You can use either

- "Ellipsoidal limits" (T^2 -based limits)
 $\mu_i - \mu_j = \bar{x}_i - \bar{x}_j \pm K_\alpha \sqrt{\{(1/n)(s_{ii} - 2s_{ij} + s_{jj})\}}$

with $K_\alpha = \sqrt{\{\chi_q^2(\alpha)\}}$ (large sample) or $K_\alpha = \sqrt{\{q \times f_e F_{q, f_e - q + 1}(\alpha)\}}$ (small sample normal).

or

- Limits based on Bonferroni t or z
 $\mu_i - \mu_j = \bar{x}_i - \bar{x}_j \pm \tilde{K}_\alpha \sqrt{\{(1/n)(s_{ii} - 2s_{ij} + s_{jj})\}}$
 $\tilde{K}_\alpha = t_{n-1}((\alpha/M)/2)$ or $\tilde{K}_\alpha = z((\alpha/M)/2)$.

with Bonferroni factor $M = p(p-1)/2$.

Comment: When p is large, M can be very large.

Suppose you are interested in M specific linear combinations $\mathbf{l}_j'\theta$, $j = 1, \dots, M$.

You can estimate each $\mathbf{l}_j'\theta$ by $\mathbf{l}_j'\hat{\theta}$ with estimated standard error

$$\hat{\sigma}_{\mathbf{l}_j'\hat{\theta}} = \sqrt{\{\mathbf{l}_j'\hat{V}[\hat{\theta}]\mathbf{l}_j\}}, j = 1, \dots, M$$

Example: In a repeated measures situation with mean vector μ , you might want to compare all $M = p(p-1)/2$ pairs μ_i and μ_j . That is, you are interested in all these $p(p-1)/2$ linear combinations

$$\mu_1 - \mu_2, \mu_1 - \mu_3, \dots, \mu_1 - \mu_p, \mu_2 - \mu_3, \dots, \mu_2 - \mu_p, \dots, \mu_{p-1} - \mu_p$$

The $p(p-1)/2$ associated \mathbf{l}_{ij} 's are

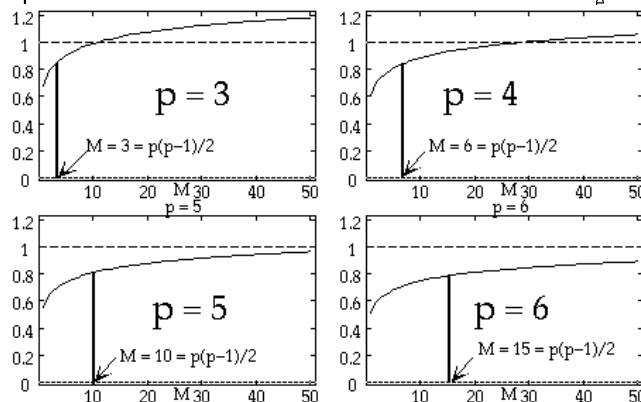
$$\mathbf{l}_{12} = [1, -1, 0, \dots, 0]', \mathbf{l}_{13} = [1, 0, -1, \dots, 0]', \dots, \mathbf{l}_{1p} = [1, 0, \dots, 0, -1]', \mathbf{l}_{23} = [0, 1, -1, \dots, 0]', \dots, \mathbf{l}_{2p} = [0, 1, \dots, 0, -1]', \dots, \mathbf{l}_{p-1,p} = [0, 0, 0, \dots, 0, 1, -1]' \text{ (contrasts)}$$

Note: $\theta_1 = \theta_2 = \dots = \theta_q \Leftrightarrow \mathbf{l}_{jk}'\theta = 0$, all $j < k$

Which intervals are shorter? Apparently, for $M=p(p-1)/2$, always the Bonferroni t or z. Here are plots against M of ratios

$$\tilde{K}_\alpha / K_\alpha = \frac{t_{f_e}(.025/M)}{\sqrt{\{(p \times f_e / (f_e - p + 1)) F_{p, f_e - p + 1}(.05)\}}}$$

for $p = 3, 4, 5, 6$ ($M=3, 6, 10, 15$), $f_e = 50$



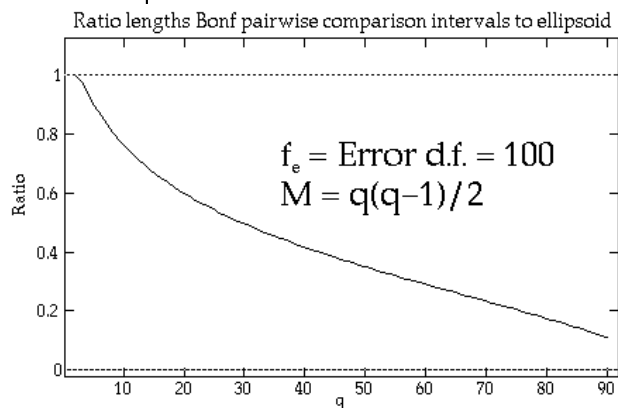
Ratio < 1 means Bonferroni intervals are shorter. For $p = 3$, only when $M > 12$ are ellipsoidal limits shorter. For $p = 6$, even with $M = 50$, Bonferroni limits are substantially shorter.

When all \mathbf{l}_j 's are *contrasts*, that is $\sum_{1 \leq k \leq q} \mathbf{l}_{kj} = 0$, you get slightly shorter ellipsoidal limits, by replacing q by $q - 1$, that is using

$$K'_\alpha = \sqrt{\{(q-1) \times f_e F_{q-1, f_e-q+2}(\alpha) / (f_e - q + 2)\}}$$

Here $f_e - q + 2 = f_e - (q-1) + 1$

Here is a plot against number of parameters q of ratio of interval lengths Bonferroniized by $M = q(q-1)/2$ to these shorter ellipsoidal limits,



As the dimension goes up, Bonferroni limits improve relative to the ellipsoidal limits.

Conclusion: *Never*, except possibly for very large M , use the ellipsoidal limits for a set of M linear combinations or comparisons that has been selected before seeing the data. When M is large, use ellipsoidal limits only when $\tilde{K}_\alpha / K'_\alpha > 1$.

Ellipsoidal limits have one advantage: They can be used with any \mathbf{l} , including an \mathbf{l} selected *after* seeing the data. This is because they apply to *all* \mathbf{l} simultaneously.

The ellipsoidal limits are similar to Sheffe multiple comparison limits.