

Displays for Statistics 5401

Lecture 4

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Class Web Page

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Matrix of 1's is useful

$$\mathbf{1}_m = \begin{bmatrix} 1 \\ 1 \\ 1 \\ \dots \\ 1 \end{bmatrix}, m \times 1 \text{ a column vector of } m \text{ 1's}$$

Using $\mathbf{1}_n$ you can use matrix multiplication to express a sum $\sum_{1 \leq j \leq n} x_j$:

Suppose $\mathbf{x} = [x_1, x_2, \dots, x_n]'$, then

$$\mathbf{1}_n' \mathbf{x} = \sum_{1 \leq j \leq n} 1 \times x_j = \sum_{1 \leq j \leq n} x_j.$$

You generate $\mathbf{1}_n$ in MacAnova by `rep(1,n)`:

```
Cmd> a <- matrix(vector(1.04,0.696, -0.651,0.13, 1.5,1.61), 2)
Cmd> a # m = 2 rows, n = 3 columns
(1,1)      1.04      -0.651      1.5
(2,1)      0.696      0.13      1.61
Cmd> sum(a) # black box column sums
(1,1)      1.736      -0.521      3.11
Cmd> ones_2 <- rep(1,2)
Cmd> ones_2' %*% a # white box column sums
(1,1)      1.736      -0.521      3.11
```

Or by a simple macro

```
Cmd> ones <- macro("rep(1,$1)"); ones(5) # same as rep(1,5)
(1)      1      1      1      1      1
Cmd> ones(2)' %*% a # exactly same as rep(2,1)' %*% a
(1,1)      1.736      -0.521      3.11
```

Rank of a matrix

Let $\mathbf{A} = [\mathbf{A}_1, \mathbf{A}_2, \dots, \mathbf{A}_n] = [\mathbf{a}_1, \mathbf{a}_2, \dots, \mathbf{a}_m]'$
 (by columns \mathbf{A}_j) (by rows \mathbf{a}_i')

be a m by n matrix.

- Columns \mathbf{A}_j are m by 1
- Rows \mathbf{a}_i' are 1 by n

Also let \mathbf{e}_k^ℓ be column k of \mathbf{I}_ℓ , that is

$$\mathbf{e}_k^\ell = \begin{bmatrix} 0 \\ \dots \\ 0 \\ 1 \\ 0 \\ \dots \\ 0 \end{bmatrix} \begin{matrix} 1 \\ \dots \\ k-1 \\ k \\ k+1 \\ \dots \\ \ell \end{matrix} \quad \begin{matrix} \text{length } \ell \\ \text{column vector} \end{matrix}$$

Example: $\mathbf{I}_4 = \begin{bmatrix} 1 & 0 & \underline{0} & 0 \\ 0 & 1 & \underline{0} & 0 \\ 0 & 0 & \underline{1} & 0 \\ 0 & 0 & \underline{0} & 1 \end{bmatrix}$, so $\mathbf{e}_3^4 = \begin{bmatrix} 0 \\ 0 \\ 1 \\ 0 \end{bmatrix}$

Fact: For any m by n \mathbf{A} ,

$\mathbf{A} = \sum_{1 \leq j \leq n} \mathbf{A}_j (\mathbf{e}_j^n)'$, sum of n outer products
 and

$\mathbf{A} = \sum_{1 \leq i \leq m} \mathbf{e}_i^m \mathbf{a}_i'$, sum of m outer products.

Conclusion:

- You can decompose *any* $m \times n$ matrix as a sum of outer products and you never need more than $\min(m,n)$ outer products to do it.
- Such a decomposition is not unique.

Vocabulary

The *rank* of \mathbf{A} is the smallest number of non-zero outer products needed to represent \mathbf{A} as a sum of outer products.

It should be clear that

- $\text{rank}(\mathbf{A}) \leq \min(m,n)$
 since $\min(m,n)$ always suffices.

Particular Cases

- A column vector $\mathbf{x} \neq \mathbf{0}$ has rank 1
- A row vector $\mathbf{x}' \neq \mathbf{0}$ has rank 1
- When $\mathbf{x} \neq \mathbf{0}$ and $\mathbf{y} \neq \mathbf{0}$ are vectors, then \mathbf{xy}' has rank 1 (it's a single outer product).
- rank of $\mathbf{D} = \text{diag}(d_1, d_2, \dots, d_p)$
 = number of $d_j \neq 0$

$$\text{rank}\left(\begin{bmatrix} 3 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & -2 \end{bmatrix}\right) = 2$$
- rank of $\mathbf{I}_p = p$

Vocabulary

A collection $\{\mathbf{l}_1, \mathbf{l}_2, \dots, \mathbf{l}_s\}$ of non-zero vectors is *linearly independent* when, for each \mathbf{l}_j , it's impossible to find c_k 's to express \mathbf{l}_j in terms of the other \mathbf{l}_k 's as $\mathbf{l}_j = \sum_{k \neq j} c_k \mathbf{l}_k$

Facts

When $\mathbf{A} = \sum_{1 \leq j \leq s} \mathbf{l}_j \mathbf{r}_j'$ has rank s , then

- $\mathbf{l}_j \neq \mathbf{0}, \mathbf{r}_j \neq \mathbf{0}, j = 1, \dots, s$
- $\{\mathbf{l}_1, \mathbf{l}_2, \dots, \mathbf{l}_s\}$ are linearly independent
- $\{\mathbf{r}_1, \mathbf{r}_2, \dots, \mathbf{r}_s\}$ are linearly independent

Conversely, when $\{\mathbf{l}_1, \mathbf{l}_2, \dots, \mathbf{l}_s\}$ and $\{\mathbf{r}_1, \mathbf{r}_2, \dots, \mathbf{r}_s\}$ are linearly independent, then $\mathbf{A} = \sum_{1 \leq j \leq s} \mathbf{l}_j \mathbf{r}_j'$ has rank s

An *important* consequence:

When $\{\mathbf{l}_j\}$ and $\{\mathbf{r}_j\}$ are linearly independent

$\mathbf{B} = \sum_{1 \leq j \leq s} \lambda_j \mathbf{l}_j \mathbf{r}_j'$ has rank $s^* < s$ if and only if $\lambda_j \neq 0$ for exactly s^* of the λ 's

Vocabulary

$m \times n$ matrix \mathbf{A} has *full rank* when $\text{rank}(\mathbf{A}) = \min(m,n)$

Determinant of a matrix

If $\mathbf{A} = [a_{ij}]$ is a $p \times p$ (square) matrix, its *determinant* is

$$\det(\mathbf{A}) = \sum_{\{j_1 j_2 \dots j_p\}} \pm a_{1j_1} a_{2j_2} \dots a_{pj_p},$$

a sum of $p! = p \times (p-1) \times \dots \times 3 \times 2 \times 1$ products. Each product has one element from each row and from each column.

Sometimes $\det(\mathbf{A})$ is notated $|\mathbf{A}|$.

- When $p = 2$, $\det(\mathbf{A}) = a_{11}a_{22} - a_{12}a_{21}$

For any p ,

- $\det(\text{diag}[a_1, a_2, \dots, a_p]) = a_1 \times a_2 \times \dots \times a_p$
- $\det(\mathbf{I}_m) = 1 \times 1 \times \dots \times 1 = 1$
- $\det(\mathbf{AB}) = \det(\mathbf{BA}) = \det(\mathbf{A}) \times \det(\mathbf{B})$, when \mathbf{A} and \mathbf{B} are square and the same size
- $\det(\mathbf{A}') = \det(\mathbf{A})$ (determinant of transpose = determinant of matrix)

In MacAnova use `det(a)`.

```
Cmd> a <- matrix(vector(17,3, 2,-1),2); a
(1,1)      17      2      2 by 2 matrix
(2,1)      3      -1      det = a11 a22 - a12 a21

Cmd> det(a)
(1)      -23

Cmd> a[1,1]*a[2,2] - a[1,2]*a[2,1] # a_11*a_22 - a_12*a_21
(1,1)      -23
```

Trace of a matrix

The *trace* of a square matrix is the sum of its diagonal elements

$$\text{tr}(\mathbf{A}) = \text{trace}(\mathbf{A}) \equiv \sum_{1 \leq i \leq p} a_{ii}$$

- $p = 2$: $\text{trace}(\mathbf{A}) = a_{11} + a_{22}$
- $p = 3$: $\text{trace}(\mathbf{A}) = a_{11} + a_{22} + a_{33}$
- When \mathbf{A} and \mathbf{B} are the same size, $\text{trace}(\mathbf{A} + \mathbf{B}) = \text{trace}(\mathbf{A}) + \text{trace}(\mathbf{B})$
- When \mathbf{A} is $p \times q$ and \mathbf{B} $q \times p$, $\mathbf{A B}$ and $\mathbf{B A}$ are defined and square and $\text{trace}(\mathbf{A B}) = \text{trace}(\mathbf{B A})$
- $\text{trace}(\mathbf{A}'\mathbf{A}) = \text{trace}(\mathbf{AA}')$ = $\sum_i \sum_j a_{ij}^2$, the sum of squares of all the elements of \mathbf{A}

The last is useful if you are trying to find a matrix $\hat{\mathbf{A}}$ of some type that is close to \mathbf{A} in the least squares sense, that is you want to minimize

$$\sum_i \sum_j (\hat{a}_{ij} - a_{ij})^2 = \text{trace}((\hat{\mathbf{A}} - \mathbf{A})'(\hat{\mathbf{A}} - \mathbf{A}))$$

MacAnova: Use `trace(a)`.

```
Cmd> deviations # previously defined matrix
(1,1)    -0.366    0.421    0.336    -0.764
(2,1)    -1.407    0.284    0.919    0.762
(3,1)     0.489   -0.173   -1.039    2.405

Cmd> trace(deviations' **% deviations)
(1)      11.626

Cmd> trace(deviations **% deviations')
(1)      11.626

Cmd> sum(vector(deviations)^2) #
(1)      11.626      Sum of squares of elements
```

`vector(deviations)` unravels the matrix `deviations` into a long vector.

```
Cmd> trace(deviations) # illegal
ERROR: argument to trace() not REAL square matrix
```

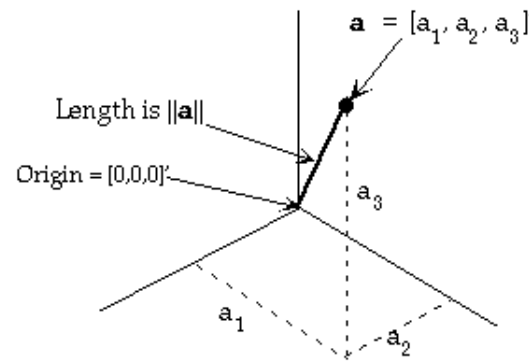
You could also use `sum(diag(a))` instead of `trace(a)`.

Some vocabulary and facts

The *length* or *norm* of a (column) vector $\mathbf{a} = [a_1 \ a_2 \ \dots \ a_p]'$:

$$\|\mathbf{a}\| = \sqrt{(\mathbf{a}'\mathbf{a})} = \sqrt{\{\sum_{1 \leq i \leq p} a_i^2\}}$$

If \mathbf{a} represents a point in p -dimensional space, $\|\mathbf{a}\|$ is the Euclidean distance of \mathbf{a} from the origin $\mathbf{0} = [0,0,\dots,0]'$.



$p = 3$

$\|\mathbf{a}\|$ is length of heavy line

MacAnova "Length" here is different from `length(a) = number of elements in a`.

Simple macro to compute $\|\mathbf{a}\|$

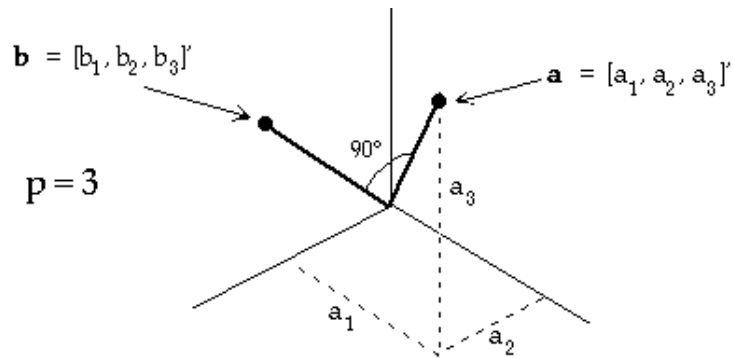
```
Cmd> norm <- macro("sqrt(sum(($1)^2))")
Cmd> z <- vector(1.1, -2.3, 4.5)
Cmd> norm(z) # compute sqrt((1.1)^2 + (-2.3)^2 + (4.5)^2)
(1)      5.172
```

Vocabulary

Vectors $\mathbf{a} \neq \mathbf{0}$ and $\mathbf{b} \neq \mathbf{0}$ with the same dimension p are *orthogonal* (*perpendicular*) when their inner product is 0:

$$\mathbf{a}'\mathbf{b} = \sum_{1 \leq i \leq p} a_i b_i = 0.$$

If \mathbf{a} and \mathbf{b} represent points in p -dimensional space, the lines from the origin to \mathbf{a} and to \mathbf{b} are perpendicular (at 90°).



The *angle* θ between two vectors \mathbf{a} and \mathbf{b} is defined by

$$\begin{aligned} \cos \theta &= \mathbf{a}'\mathbf{b} / (\|\mathbf{a}\| \times \|\mathbf{b}\|) \\ &= \sum_i a_i b_i / \sqrt{(\sum_i a_i^2 \sum_j b_j^2)} \end{aligned}$$

So $\mathbf{a}'\mathbf{b} = 0 \Rightarrow \cos \theta = 0 \Rightarrow \theta = \pm 90^\circ$

Suppose $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_r$ with all $\mathbf{u}_i \neq \mathbf{0}$, are *mutually orthogonal* ($\mathbf{u}_i' \mathbf{u}_j = 0, i \neq j$).

Then these **facts** are true:

- $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_r$ are linearly independent
- $\mathbf{U} = [\mathbf{u}_1, \dots, \mathbf{u}_r]$ ($p \times r$) has rank r .
- $\mathbf{U}'\mathbf{U} = \text{diag}(\|\mathbf{u}_1\|^2, \dots, \|\mathbf{u}_r\|^2)$ is diagonal

This last is really the definition of mutual orthogonality.

Angles and correlation coefficients

The sample correlation between two variables $\{x_i\}_{1 \leq i \leq n}$ and $\{y_i\}_{1 \leq i \leq n}$ is

$$r_{xy} = \frac{\sum_{1 \leq i \leq n} (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\{\sum_{1 \leq i \leq n} (x_i - \bar{x})^2 \sum_{1 \leq i \leq n} (y_i - \bar{y})^2\}}} = \cos \theta_{xy}$$

You can interpret $-1 \leq r_{xy} \leq 1$ as the cosine of the angle θ_{xy} between the n-vectors of deviations from the mean.

$$\tilde{X} = [x_1 - \bar{x}, \dots, x_n - \bar{x}]'$$

and

$$\tilde{Y} = [y_1 - \bar{y}, \dots, y_n - \bar{y}]'$$

When $r_{xy} \approx 1$, $\theta_{xy} \approx 0$ so \tilde{X} and \tilde{Y} point in almost the *same* direction.

When $r_{xy} \approx -1$, $\theta_{xy} \approx \pm 180^\circ$ so \tilde{X} and \tilde{Y} point in almost the *opposite* direction.

Inverse of a Matrix

Vocabulary:

The *inverse* of a p by p *square* matrix **A** is the matrix **A**⁻¹ (if one exists) such that

- **AA**⁻¹ = **I**_p = **A**⁻¹**A**.

There is *at most* one such matrix.

Example:

$$A = \begin{bmatrix} 7 & 2 \\ 4 & 4 \end{bmatrix}, \quad A^{-1} = \begin{bmatrix} 1/5 & -1/10 \\ -1/5 & 7/20 \end{bmatrix}$$

```
Cmd> a <- matrix(vector(7,4,2,4),2);a
(1,1)      7      2
(2,1)      4      4

Cmd> ainv <- matrix(vector(1/5,-1/5, -1/10,7/20),2); ainv
(1,1)      0.2     -0.1
(2,1)     -0.2     0.35

Cmd> a %*% ainv # 2 by 2 identity
(1,1)      1 -1.1102e-16   -1.1102e-16 ~ 0
(2,1)      0      1

Cmd> ainv %*% a # 2 by 2 identity
(1,1)      1      0
(2,1) -2.2204e-16   1

Cmd> solve(a) # solve() computes inverse
(1,1)      0.2     -0.1
(2,1)     -0.2     0.35
```

Note a %*% ainv and ainv %*% a aren't *exactly* **I**₂ because of rounding error.

Here are two matrices with *no* inverse:

$$\mathbf{B} = \begin{bmatrix} 1 & 3 \\ 2 & 6 \end{bmatrix}, \quad \mathbf{C} = \begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix}$$

```
Cmd> b <- matrix(vector(1,3, 2,6),2); b
(1,1)      1      2
(2,1)      3      6

Cmd> solve(b) # try to find inverse
ERROR: argument to solve() is apparently singular

Cmd> c <- matrix(vector(1,0, 0,0),2); c
(1,1)      1      0
(2,1)      0      0

Cmd> solve(c) # try to find inverse
ERROR: argument to solve() is singular
```

Vocabulary

- When \mathbf{A}^{-1} exists, \mathbf{A} is *invertible* or *non-singular*
- \mathbf{A}^{-1} does not exist $\Rightarrow \mathbf{A}$ is *singular*

More "facts" when \mathbf{A} is non-singular:

- $\mathbf{A}^{-1}\mathbf{A} = \mathbf{AA}^{-1} = \mathbf{I}_p$ (definition of \mathbf{A}^{-1})
- $(\mathbf{A}^{-1})^{-1} = \mathbf{A}$ (inverse of \mathbf{A}^{-1} is \mathbf{A})
- $(\mathbf{A}')^{-1} = (\mathbf{A}^{-1})'$ (transpose of inverse is inverse of transpose)
- \mathbf{A} and \mathbf{B} non-singular $\Rightarrow (\mathbf{AB})^{-1} = \mathbf{B}^{-1}\mathbf{A}^{-1}$

Using vectors and matrices to represent data

Univariate Data ($p = 1$)

A *univariate* data set consists of n observations x_1, \dots, x_n on $p = 1$ variable.

You represent it by the column vector of length n ($n \times 1$ matrix)

$$\mathbf{X} = \begin{bmatrix} x_1 \\ x_2 \\ \dots \\ x_n \end{bmatrix} = [x_1, x_2, \dots, x_n]', \quad n \text{ by } 1$$

You *could* represent the data by the *row* vector

$$[x_1, x_2, \dots, x_n].$$

but that's not the convention we use.

We use the column vector form *only*.

The sum of the data is $\mathbf{1}'_n \mathbf{X} = \sum_{1 \leq i \leq n} x_i$.

Multivariate Data ($p > 1$)

Suppose your data are n *multivariate* observations $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n$ (p by 1 vectors), with

$$\mathbf{x}_i = [x_{i1}, x_{i2}, \dots, x_{ip}]'$$

You can represent all the data by the $n \times p$ **data matrix**

$$\mathbf{X} = \begin{bmatrix} \mathbf{x}_1' \\ \mathbf{x}_2' \\ \dots \\ \mathbf{x}_n' \end{bmatrix} = [\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_p] = [x_{ij}]_{1 \leq i \leq n, 1 \leq j \leq p}$$

- Column vector \mathbf{X}_j = all the data on *variable* j .
- Row vector \mathbf{x}_i' = $[x_{i1}, x_{i2}, \dots, x_{ip}]$ = all the data on *case* i , expressed as the transpose of the column vector \mathbf{x}_i .

Sums of squares and products

Suppose $\mathbf{X} = [\mathbf{X}_1, \dots, \mathbf{X}_p]$ is n by p , then

- The *diagonal elements* of $\mathbf{X}'\mathbf{X} = [X_{jk}]_{1 \leq j \leq p, 1 \leq k \leq p}$ are *sums of squares* $\sum_{1 \leq i \leq n} x_{ij}^2 = \mathbf{X}_j' \mathbf{X}_j$
- The *off diagonal elements* of $\mathbf{X}'\mathbf{X}$ are *sums of products* $\sum_{1 \leq i \leq n} x_{ij} x_{ik} = \mathbf{X}_j' \mathbf{X}_k, j \neq k$

$\mathbf{X}'\mathbf{X}$ is also a *sum of outer products* $\mathbf{x}_i \mathbf{x}_i'$

$$\mathbf{X}'\mathbf{X} = \sum_{1 \leq i \leq n} \mathbf{x}_i \mathbf{x}_i', \mathbf{x}_i' \text{ a row of } \mathbf{X}.$$

When the data are univariate, \mathbf{X} is n by 1 and $\mathbf{X}'\mathbf{X} = \sum_{1 \leq i \leq n} x_i^2$.

Important mnemonic

A square x_i^2 in a univariate formula often becomes an outer product $\mathbf{x}_i \mathbf{x}_i'$ in a related multivariate formula.

$$\sum_{1 \leq i \leq n} x_i^2 \Rightarrow \sum_{1 \leq i \leq n} \mathbf{x}_i \mathbf{x}_i'$$

MacAnova

```

Cmd> x <- matrix(vector(2.4,12.3,10.6,15.1,-1.3, \
  22.9,15.7,15.7,17.2,22.5, 44,32.7,35.2,33.5,26.7), 5)

Cmd> x # n = 5; p = 3
(1,1)      2.4      22.9      44
(2,1)     12.3     15.7     32.7
(3,1)     10.6     15.7     35.2
(4,1)     15.1     17.2     33.5
(5,1)     -1.3     22.5     26.7

```

$$\mathbf{x}$$

```

Cmd> xx <- x' %*% x; xx # p by p (3 by 3)
(1,1)     499.11     644.96     1352.1
(2,1)     644.96     1819.5     3250.6
(3,1)     1352.1     3250.6     6079.5

```

$$\mathbf{x}'\mathbf{x}$$

```

Cmd> sum(x[,1]*x[,3]) # summed products of cols 1 and 3
(1,1)     1352.1     xx[1,3]

```

Use loop to compute $\mathbf{X}'\mathbf{X}$ as a sum of outer products $\sum_{1 \leq i \leq n} \mathbf{x}_i \mathbf{x}_i'$:

```

Cmd> xx1 <- dmat(3,0) # 3 by 3 starting matrix of 0's
Cmd> n <- nrows(x) # sample size
Cmd> for(i,1,n){ # X'X as sum of outer products
  @xi <- vector(x[i,]) # column i of x
  xx1 <- xx1 + outer(@xi,@xi)
};} # ";" prevents extraneous output

Cmd> xx1 # same as xx
(1,1)     499.11     644.96     1352.1
(2,1)     644.96     1819.5     3250.6
(3,1)     1352.1     3250.6     6079.5

```

MacAnova

The ";" before the final "}" prevents printing each time through the loop. @xi is a temporary variable that is automatically deleted after the loop.

Statistical application of matrix formulas

Suppose $\mathbf{X} = \begin{bmatrix} \mathbf{x}_1' \\ \mathbf{x}_2' \\ \dots \\ \mathbf{x}_n' \end{bmatrix} = [\mathbf{X}_1 \ \mathbf{X}_2 \ \dots \ \mathbf{X}_p]$ is a

data matrix containing n cases of p variables.

The *sample mean vector* is

$$\bar{\mathbf{x}} = (1/n) \sum_{1 \leq i \leq n} \mathbf{x}_i = \begin{bmatrix} \bar{X}_1 \\ \bar{X}_2 \\ \dots \\ \bar{X}_p \end{bmatrix},$$

where

$\bar{x}_j = (1/n) \sum_{1 \leq i \leq n} x_{ij} = (1/n) \mathbf{1}_n' \mathbf{X}_j$ $j = 1, \dots, p$, is a *univariate average*.

The *row vector* $\bar{\mathbf{x}}' = (1/n)\mathbf{1}'\mathbf{X}$

```
Cmd> xbar <- rep(1,n)' %*% x / n
```

This gives the same result as $\text{sum}(\mathbf{x})/n$

```
Cmd> equal(xbar,sum(x)/n)
(1) T
```

The **sample variance** (or **covariance** or **variance/covariance**) **matrix** is

$$\mathbf{S} = [s_{ij}] \equiv (n-1)^{-1} \sum_{1 \leq i \leq n} (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})'$$

Compare this with the univariate sample variance $s^2 = (n-1)^{-1} \sum_{1 \leq i \leq n} (x_i - \bar{x})^2$.

If $\tilde{\mathbf{X}} = [\tilde{\mathbf{x}}_1, \tilde{\mathbf{x}}_2, \dots, \tilde{\mathbf{x}}_n]'$, $\tilde{\mathbf{x}}_i = \mathbf{x}_i - \bar{\mathbf{x}}$, the matrix of deviation of the observations from their sample mean

$$\mathbf{S} = (n-1)^{-1} \tilde{\mathbf{X}}' \tilde{\mathbf{X}} = (n-1)^{-1} \sum_{1 \leq i \leq n} \tilde{\mathbf{x}}_i \tilde{\mathbf{x}}_i'$$

Note: This differs from a similar definition with a divisor of n .

$$\mathbf{S}_n = \sum (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})' / n$$

As estimators of the population variance matrix Σ (not yet defined), \mathbf{S} is unbiased and \mathbf{S}_n is biased.

- On the *diagonal*, $s_{jj} = (n-1)^{-1} \sum (x_{ij} - \bar{x}_j)^2$ are the usual *sample variances* s_j^2 .
 $\sqrt{s_{ii}}$ = *sample standard deviation*
- The *off-diagonal* elements
 $s_{jk} = (n-1)^{-1} \sum (x_{ij} - \bar{x}_j)(x_{ik} - \bar{x}_k)$
are the *sample covariances*.

The divisor $n-1$ is the *degrees of freedom*. The n deviations $\mathbf{x}_i - \bar{\mathbf{x}}$ from the mean satisfy one linear equality, namely, $\sum_{1 \leq i \leq n} (\mathbf{x}_i - \bar{\mathbf{x}}) = 0$

Important observation:

You can get the multivariate ($p > 1$) formula

$$\mathbf{S} = (n-1)^{-1} \sum_{1 \leq i \leq n} (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})'$$

the univariate ($p = 1$) formula

$$s^2 = (n-1)^{-1} \sum_{1 \leq i \leq n} (x_i - \bar{x})^2$$

Replace $(x_i - \bar{x})^2$ by $(\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})'$.